# Classes of Ordinary Differential Equations Obtained for the Probability Functions of Exponential and Truncated Exponential Distributions

Hilary I. Okagbue, Pelumi E. Oguntunde, Member, AENG, Abiodun A. Opanuga,

Enahoro A. Owoloko

Abstract— In this paper, the differential calculus was used to obtain some classes of ordinary differential equations (ODE) for the probability density function, quantile function, survival function, inverse survival function and reversed hazard function of the exponential and truncated exponential distributions. However, the ODE of the hazard function of the exponential distribution cannot be obtained because of the constant failure nature of the distribution. The ODE of the truncated exponential distribution was obtained. The stated necessary conditions required for the existence of the ODEs are consistent with the various parameters that defined the distributions. Solutions of these ODEs by using numerous available methods are new ways of understanding the nature of the probability functions that characterize the distribution.

*Index Terms*— Exponential distribution, differential calculus, probability density function, survival function, quantile function, reversed hazard function.

#### I. INTRODUCTION

EXPONENTIAL distribution is used to model the time between events or events between a time intervals. Because of the flexibility of the distribution and constant hazard, the distribution has seen many modifications in form of compounding, exponentiation, transmuted and so on. Some are distributions include: multivariate exponential distribution [1], generalized exponential distribution [2], beta exponential distribution [3], beta generalized exponential distribution [4], quadratic exponential binary distribution [5], generalized exponential Poisson distribution [6], extended Poisson exponential distribution [7], gamma exponentiated exponential distribution [8], convoluted betaexponential distribution [9], weighted exponential distribution [10], fractional beta exponential distribution [11], exponential Poisson Logarithmic distribution [12], exponentiated Kumaraswamy-exponential distribution [13], Weibull exponential distribution [14], Lindley exponential distribution [15], weighted exponential distribution [16].

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H. I. Okagbue, P. E. Oguntunde, A. A. Opanuga and E.A. Owoloko are with the Department of Mathematics, Covenant University, Ota, Nigeria.

hilary.okagbue@covenantuniversity.edu.ng pelumi.oguntunde@covenantuniversity.edu.ng abiodun.opanuga@covenantuniversity.edu.ng alfred.owoloko@covenantuniversity.edu.ng The truncated exponential distribution is a sub model of the parent exponential distribution obtained by the restriction of some selected domain of exponential distribution. The details can be obtained from the notes of [17]. Lominashvili and Patsatsia [18] obtained the estimates of maximum likelihood of the distribution. Korkmaz and Genç [19]observed that their proposed distribution known as the two-sided generalized exponential distribution is the mixture of truncated exponential distribution and truncated generalized exponential distribution earlier proposed by [20].

Exponential has been recently used to model the following: analysis of plates in solid dynamics [21], infection rate [22] and ecology [23].

The aim of this research is to develop ordinary differential equations (ODE) for the probability density function (PDF), Quantile function (QF), survival function (SF), inverse survival function (ISF) and reversed hazard function (RHF) of exponential and truncated exponential distributions by the use of differential calculus. Also the ODE of hazard function of the truncated exponential distribution was also obtained. Calculus is a very key tool in the determination of mode of a given probability distribution and in estimation of parameters of probability distributions, amongst other uses. The research is an extension of the ODE to other probability functions other than the PDF. Similar works done where the PDF of probability distributions was expressed as ODE whose solution is the PDF are available. They include: Laplace distribution [24], beta distribution [25], raised cosine distribution [26], Lomax distribution [27], beta prime distribution or inverted beta distribution [28].

# I. EXPONENTIAL DISTRIBUTION PROBABILITY DENSITY FUNCTION

The probability density function of the Cauchy distribution is given by;

The probability density function (PDF) of the exponential distribution is given by;

$$f(x) = \lambda e^{-\lambda x} \tag{1}$$

The first order differential equation for the PDF can be obtained from the differentiation of equation (1);

$$f'(x) = -\lambda^2 e^{-\lambda x}$$
<sup>(2)</sup>

The condition necessary for the existence of the equation is

$$\lambda > 0, x \ge 0$$
  
$$f'(x) = -\lambda(\lambda e^{-\lambda x}) = -\lambda f(x)$$
(3)

The first order ordinary differential for the probability density function of the exponential distribution is given as;

$$f'(x) + \lambda f(x) = 0 \tag{4}$$

$$f(0) = \lambda \tag{5}$$

To obtain the second order differential equation, differentiate equation (2) to obtain;

$$f''(x) = \lambda^3 e^{-\lambda x} \tag{6}$$

The condition necessary for the existence of the equation is  $\lambda > 0, x \ge 0$ .

$$f''(x) = -\lambda(-\lambda^2 e^{-\lambda x}) = -\lambda f'(x)$$
(7)

The second order ordinary differential for the probability density function of the exponential distribution is given as;

$$f''(x) + \lambda f'(x) = 0 \tag{8}$$

$$f'(0) = -\lambda^2 \tag{9}$$

To obtain the third order differential equation, differentiate equation (6) to obtain;

$$f'''(x) = -\lambda^4 e^{-\lambda x} \tag{10}$$

The condition necessary for the existence of the equation is  $\lambda > 0, x \ge 0$ .

$$f'''(x) = -\lambda(\lambda^3 e^{-\lambda x}) = -\lambda f''(x)$$
(11)

The third order ordinary differential for the probability density function of the exponential distribution is given as;

$$f'''(x) + \lambda f''(x) = 0$$
(12)

$$f''(0) = \lambda^3 \tag{13}$$

Similarly, higher order ordinary differential equations can be obtained such as;

$$f''(x) + \lambda f'''(x) = 0$$
(14)

$$f^{\nu}(x) + \lambda f^{\prime\nu}(x) = 0 \tag{15}$$

#### **QUANTILE FUNCTION**

The Quantile function (QF) of the exponential distribution is given by;

$$Q(p) = -\frac{\ln(1-p)}{\lambda} \tag{16}$$

The first order differential equation for the QF can be obtained from the differentiation of equation (16);

$$Q'(p) = \frac{1}{\lambda} (1-p)^{-1}$$
(17)

Equation (17) can also be written as;

$$Q'(p) = \frac{1}{\lambda(1-p)} \tag{18}$$

The condition necessary for the existence of the equation is  $\lambda > 0, 0 .$ 

Simplify equation (18) to obtain;  

$$\lambda(1-p)Q'(p) = 1$$
 (19)

The first order ordinary differential for the Quantile function of the exponential distribution is given as;

$$\lambda(1-p)Q'(p) - 1 = 0 \tag{20}$$

$$Q(0.1) = \frac{0.1054}{\lambda} \tag{21}$$

To obtain the second order differential equation, differentiate equation (17) to obtain;

$$Q''(p) = \frac{1}{\lambda} (1-p)^{-2}$$
(22)

Equation (22) can also be written as;

$$Q''(p) = \frac{1}{\lambda(1-p)^2}$$
(23)

The condition necessary for the existence of the equation is  $\lambda > 0, 0 .$ 

Simplify equation (23) to obtain;

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$$(1-p)\left\lfloor\frac{1}{\lambda(1-p)^2}\right\rfloor = \frac{1}{\lambda(1-p)}$$
(24)

$$(1-p)Q''(p) = Q'(p)$$
 (25)

The second order ordinary differential for the Quantile function of the exponential distribution is given as;

$$(1-p)Q''(p) - Q'(p) = 0$$
(26)

$$Q'(0.1) = \frac{10}{9\lambda}$$
(27)

To obtain the third order differential equation, differentiate equation (22) to obtain;

$$Q'''(p) = \frac{2}{\lambda} (1-p)^{-3}$$
(28)

Equation (28) can also be written as;

$$Q'''(p) = \frac{2}{\lambda(1-p)^3}$$
(29)

The condition necessary for the existence of the equation is  $\lambda > 0, 0 .$ 

Simplify equation (29) to obtain;

$$\frac{(1-p)}{2} \left\lfloor \frac{2}{\lambda(1-p)^3} \right\rfloor = \frac{1}{\lambda(1-p)^2}$$
(30)

$$\frac{(1-p)}{2}Q'''(p) = Q''(p)$$
(31)

The third order ordinary differential for the Quantile function of the exponential distribution is given as;

$$(1-p)Q'''(p) - 2Q''(p) = 0$$
(32)

$$Q''(0.1) = \frac{100}{81\lambda}$$
(33)

Similarly, higher order ordinary differential equations can be obtained such as;

$$(1-p)Q''(p) - 3Q'''(p) = 0$$
(34)

$$(1-p)Q^{\nu}(p) - 4Q^{\prime\nu}(p) = 0$$
(35)

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# SURVIVAL FUNCTION

The survival function (SF) of the exponential distribution is given by;

$$S(t) = e^{-\lambda t} \tag{36}$$

The first order differential equation for the SF can be obtained from the differentiation of equation (36);

$$S'(t) = -\lambda e^{-\lambda t} \tag{37}$$

The condition necessary for the existence of the equation is  $\lambda > 0, t \ge 0$ .

$$S'(t) = -\lambda(e^{-\lambda t}) = -\lambda S(t)$$
(38)

The first order ordinary differential for the survival function of the exponential distribution is given as;

$$S'(t) + \lambda S(t) = 0 \tag{39}$$

$$S(0) = 1 \tag{40}$$

To obtain the second order differential equation,

$$S''(t) = \lambda^2 e^{-\lambda t}$$
<sup>(41)</sup>

The condition necessary for the existence of the equation is  $\lambda > 0, t \ge 0$ .

$$S''(t) = -\lambda(-\lambda e^{-\lambda t}) = -\lambda S'(t)$$
(42)

The second order ordinary differential for the survival function of the exponential distribution is given as;

$$S''(t) + \lambda S'(t) = 0 \tag{43}$$

$$S'(0) = \lambda \tag{44}$$

Similarly, higher order ordinary differential equations can be obtained such as;

$$S'''(t) + \lambda S''(t) = 0 \tag{45}$$

$$S'''(t) + \lambda S'''(t) = 0$$
(46)

$$S^{\nu}(t) + \lambda S^{\prime\nu}(t) = 0 \tag{47}$$

## INVERSE SURVIVAL FUNCTION

The inverse survival function (ISF) of the exponential distribution is given by;

$$Q(p) = -\frac{\ln p}{\lambda} \tag{48}$$

The first order differential equation for the ISF can be obtained from the differentiation of equation (48);

$$Q'(p) = -\frac{1}{\lambda p} \tag{49}$$

The condition necessary for the existence of the equation is  $\lambda > 0, 0 .$ 

Simplify equation (49) to obtain;

$$\lambda p Q'(p) = -1 \tag{50}$$

The first order ordinary differential for the inverse survival function of the exponential distribution is given as;

$$\lambda p Q'(p) + 1 = 0 \tag{51}$$

$$Q(0.1) = \frac{2.3026}{\lambda} \tag{52}$$

To obtain the second order differential equation, differentiate equation (49) to obtain;

$$Q''(p) = \frac{1}{\lambda p^2}$$
(53)

The condition necessary for the existence of the equation is  $\lambda > 0, 0 .$ 

Simplify equation (53) to obtain;

$$p\left\lfloor\frac{1}{\lambda p^2}\right\rfloor = \frac{1}{\lambda p}$$
(54)

$$pQ''(p) = -Q'(p) \tag{55}$$

The second order ordinary differential for the inverse survival function of the exponential distribution is given

$$pQ''(p) + Q'(p) = 0$$
 (56)

$$Q'(0.1) = -\frac{10}{\lambda} \tag{57}$$

To obtain the third order differential equation, differentiate equation (53) to obtain;

$$Q'''(p) = -\frac{2}{\lambda p^3}$$
(58)

The condition necessary for the existence of the equation is  $\lambda > 0, 0 .$ 

Simplify equation (58) to obtain;

$$\frac{p}{2} \left[ -\frac{2}{\lambda p^3} \right] = -\frac{1}{\lambda p^2}$$
(59)

$$\frac{p}{2}Q'''(p) = -Q''(p) \tag{60}$$

The third order ordinary differential for the inverse survival function of the exponential distribution is given as;

$$pQ'''(p) + 2Q''(p) = 0 \tag{61}$$

$$Q''(0.1) = \frac{100}{\lambda} \tag{62}$$

## **REVERSED HAZARD FUNCTION**

The reversed hazard function (RHF) of the exponential distribution is given by;

$$j(t) = \frac{\lambda e^{-\lambda t}}{1 - e^{-\lambda t}}$$
(63)

The first order differential equation for the RHF can be obtained from the differentiation of equation (63);

$$j'(t) = \left\{ -\frac{\lambda e^{-\lambda t}}{e^{-\lambda t}} - \frac{\lambda e^{-\lambda t} (1 - e^{-\lambda t})^{-2}}{(1 - e^{-\lambda t})^{-1}} \right\} j(t)$$
(64)

The condition necessary for the existence of the equation is  $\lambda > 0, t \ge 0$ .

$$j'(t) = -\left(\lambda + \frac{\lambda e^{-\lambda t}}{1 - e^{-\lambda t}}\right) j(t)$$
(65)

$$j'(t) = -(\lambda + j(t))j(t) \tag{66}$$

The first order ordinary differential for the reversed hazard function of the exponential distribution is given as;

$$j'(t) + j^{2}(t) + \lambda j(t) = 0$$
(67)

$$j(1) = \frac{\lambda e^{-\lambda}}{1 - e^{-\lambda}} = \frac{\lambda}{e^{\lambda} - 1}$$
(68)

## II. TRUNCATED EXPONENTIAL DISTRIBUTION

## PROBABILITY DENSITY FUNCTION

The probability density function (PDF) of the truncated exponential distribution is given by;

$$f(x) = \frac{\frac{1}{\lambda} e^{-\frac{x}{\lambda}}}{1 - e^{-\left(\frac{b}{\lambda}\right)}}$$
(69)

The first order differential equation for the PDF can be obtained from the differentiation of equation (69);

$$f'(x) = -\frac{\frac{1}{\lambda^2} e^{-\frac{x}{\lambda}}}{1 - e^{-\left(\frac{b}{\lambda}\right)}}$$
(70)

The condition necessary for the existence of the equation is  $\lambda, b > 0, x \ge 0$ .

Simplify equation (70) using (69) to obtain;

$$f'(x) = -\frac{f(x)}{\lambda} \tag{71}$$

The first order ordinary differential for the probability density function of the truncated exponential distribution is given as;

$$\lambda f'(x) + f(x) = 0 \tag{72}$$

$$f(0) = \frac{1}{\lambda(1 - e^{-\left(\frac{b}{\lambda}\right)})}$$
(73)

To obtain the second order differential equation, differentiate equation (70) to obtain;

$$f''(x) = \frac{\frac{1}{\lambda^3} e^{-\frac{x}{\lambda}}}{1 - e^{-\left(\frac{b}{\lambda}\right)}}$$
(74)

The condition necessary for the existence of the equation is  $\lambda, b > 0, x \ge 0$ .

Simplify equation (74) to obtain;

$$f''(x) = -\frac{f'(x)}{\lambda} \tag{75}$$

The second order ordinary differential for the probability density function of the truncated exponential distribution is given as;

$$\lambda f''(x) + f'(x) = 0 \tag{76}$$

$$f'(0) = -\frac{1}{\lambda^2 (1 - e^{-\left(\frac{b}{\lambda}\right)})}$$
(77)

Similarly, higher order ordinary differential equations can be obtained such as;

$$\lambda f'''(x) + f''(x) = 0$$
(78)

$$\lambda f''(x) + f'''(x) = 0 \tag{79}$$

$$\lambda f^{\nu}(x) + f^{\prime\nu}(x) = 0 \tag{80}$$

# QUANTILE FUNCTION

The Quantile function (QF) of the truncated exponential distribution is given as;

$$Q(p) = -\lambda \ln(1 - p + p e^{-\left(\frac{b}{\lambda}\right)})$$
(81)

The first order differential equation for the QF can be obtained from the differentiation of equation (81);

$$Q'(p) = -\frac{\lambda(e^{-\frac{b}{\lambda}} - 1)}{1 - p + p e^{-\frac{b}{\lambda}}}$$
(82)

The condition necessary for the existence of the equation is  $\lambda, b > 0, 0 .$ 

Let 
$$A = e^{\frac{b}{\lambda}}$$
 (83)  
Substitute equation (83) into equation (82) to obtain;

$$Q'(p) = -\frac{\lambda(A-1)}{1-p+pA}$$

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(84) The first order ordinary differential for the Quantile function of the truncated exponential distribution is given as;  $(1 - p + pA)Q'(p) + \lambda(A - 1) = 0$ 

; 
$$(1-p+pA)Q'(p) + \lambda(A-1) = 0$$
  
(85)  $Q(0) = 0$ 

(86)

To obtain the second order differential equation, equation (82) can be written as;

$$Q'(p) = \lambda (1 - e^{-\frac{b}{\lambda}})(1 - p + p e^{-\frac{b}{\lambda}})^{-1}$$
(87)

Differentiate equation (87) to obtain; b

$$Q''(p) = \lambda (1 - e^{-\lambda})^2 (1 - p + p e^{-\lambda})^{-2}$$
(88)

The condition necessary for the existence of the equation is  $\lambda, b > 0, 0 .$ 

Re-arranging and simplifying equation (88), to obtain

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$$Q''(p) = \frac{\lambda (1 - e^{-\frac{b}{\lambda}})^2}{(1 - p + p e^{-\frac{b}{\lambda}})^2}$$
(89)

$$\lambda Q''(p) = \frac{\lambda^2 (1 - e^{-\frac{1}{\lambda}})^2}{(1 - p + p e^{-\frac{b}{\lambda}})^2}$$
(90)

$$\lambda Q''(p) = Q'^2(p) \tag{91}$$

The second order ordinary differential for the Quantile function of the truncated exponential distribution is given as;  $\lambda Q''(p) - Q'^2(p) = 0$ 

(92)  $Q'(0) = \lambda (1 - e^{-\frac{b}{\lambda}})$ 

(93) To obtain the third order differential equation, differentiate equation (88) to obtain;

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$$Q'''(p) = 2\lambda (1 - e^{-\frac{b}{\lambda}})^3 (1 - p + p e^{-\frac{b}{\lambda}})^{-3}$$
(94)

The condition necessary for the existence of the equation is  $\lambda, b > 0, 0 .$ 

Re-arranging and simplifying equation (91), to obtain

$$Q'''(p) = \frac{2\lambda (1 - e^{-\frac{b}{\lambda}})^3}{(1 - p + p e^{-\frac{b}{\lambda}})^3}$$
(95)

$$\lambda^2 Q'''(p) = \frac{2\lambda^3 (1 - e^{-\frac{b}{\lambda}})^3}{(1 - p + p e^{-\frac{b}{\lambda}})^3}$$
(96)

$$\lambda^2 Q'''(p) = Q'^3(p) \tag{97}$$

The third order ordinary differential for the Quantile function of the truncated exponential distribution is given as:  $\lambda^2 Q'''(p) - 2Q'^3(p) = 0$ 

(98) 
$$Q''(0) = \lambda (1 - e^{-\frac{b}{\lambda}})^2$$

(99) Similarly, higher order ordinary differential equations can be obtained such as;

$$\lambda^{3}Q^{\prime\nu}(p) - 3Q^{\prime4}(p) = 0 \tag{100}$$

$$\lambda^4 Q^{\nu}(p) - 4Q^{\prime 5}(p) = 0 \tag{101}$$

#### SURVIVAL FUNCTION

The survival function (SF) of the truncated exponential distribution is derived from the cumulative distribution function and is given by;

$$S(t) = \frac{e^{-\frac{t}{\lambda}} - e^{-\frac{b}{\lambda}}}{1 - e^{-\frac{b}{\lambda}}}$$
(102)

Let 
$$C = e^{\frac{b}{\lambda}}$$
 (103)

$$S(t) = \frac{e^{\frac{-\lambda}{\lambda}} - C}{1 - C} \tag{104}$$

The first order differential equation for the SF can be obtained from the differentiation of equation (104);

$$S'(t) = \frac{-\frac{1}{\lambda}e^{-\frac{t}{\lambda}}}{1-C}$$
(105)

The condition necessary for the existence of the equation is  $\lambda, b > 0, t \ge 0$ .

Some simplifications are done to reduce equation (105) to an ordinary differential equation.

From equation (104);

$$(1-C)S(t) = e^{-\frac{t}{\lambda}} - C$$
(106)

$$e^{-\lambda} = (1-C)S(t) + C$$
(107)

Substitute equation (107) into equation (105), to obtain;

$$S'(t) = -\frac{(1-C)S(t) + C}{\lambda(1-C)}$$
(108)

$$\lambda(1-C)S'(t) + (1-C)S(t) + C = 0$$
(109)

$$Let \quad D = 1 - C \tag{110}$$

The first order ordinary differential for the survival function of the truncated exponential distribution is given as;

$$\lambda DS'(t) + DS(t) + 1 - D = 0$$
(111)

$$S(0) = 1$$
 (112)

To obtain the second order differential equation, differentiate equation (105) to obtain;

$$S''(t) = \frac{\frac{1}{\lambda^2} e^{-\frac{t}{\lambda}}}{1 - C}$$
(113)

The condition necessary for the existence of the equation is  $\lambda, b > 0, t \ge 0$ .

Simplifying to obtain;

$$S''(t) = -\frac{1}{\lambda} \left[ \frac{-\frac{1}{\lambda} e^{-\frac{t}{\lambda}}}{1 - C} \right]$$
(114)

$$S''(t) = -\frac{1}{\lambda}S'(t) \tag{115}$$

The second order ordinary differential for the Survival function of the truncated exponential distribution is given as;  $\lambda S''(t) + S'(t) = 0$ 

(116) 
$$S'(0) = -\frac{1}{\lambda(1-C)}$$

(117) To obtain the third order differential equation, differentiate equation (113) to obtain;

$$S'''(t) = -\frac{\frac{1}{\lambda^3} e^{-\frac{t}{\lambda}}}{1 - C}$$
(118)

The condition necessary for the existence of the equation is  $\lambda, b > 0, t \ge 0$ .

Simplifying to obtain;

$$S'''(t) = -\frac{1}{\lambda} \left[ \frac{\frac{1}{\lambda^2} e^{-\frac{1}{\lambda}}}{1 - C} \right]$$
(119)

$$S'''(t) = -\frac{1}{\lambda}S''(t) \tag{120}$$

The third order ordinary differential for the Survival function of the truncated exponential distribution is given as;  $\lambda S'''(t) + S''(t) = 0$ 

(121) 
$$S''(0) = \frac{1}{\lambda^2 (1-C)}$$

(122) Similarly, higher order ordinary differential equations can be obtained such as;

$$\lambda S^{\prime v}(t) + S^{\prime \prime}(t) = 0 \tag{123}$$

$$\lambda S^{\nu}(t) + S^{\prime\nu}(t) = 0 \tag{124}$$

#### HAZARD FUNCTION

The hazard function (HF) of the truncated exponential distribution is the ratio of the probability density function to the survival function. This is given as;

$$h(t) = \frac{\frac{1}{\lambda} e^{-\frac{t}{\lambda}}}{e^{-\frac{t}{\lambda}} - e^{-\frac{b}{\lambda}}}$$
(125)

The first order differential equation for the HF can be obtained from the differentiation of equation (125);

$$h'(t) = \frac{\frac{1}{\lambda^2} \left( e^{-\frac{t}{\lambda}} \right)^2}{\left( e^{-\frac{t}{\lambda}} - e^{-\frac{b}{\lambda}} \right)^2} - \frac{\frac{1}{\lambda^2} \left( e^{-\frac{t}{\lambda}} \right)}{\left( e^{-\frac{t}{\lambda}} - e^{-\frac{b}{\lambda}} \right)}$$
(126)

The condition necessary for the existence of the equation is  $\lambda, b > 0, t \ge 0$ .

$$h'(t) = h^{2}(t) - \lambda h(t)$$
 (127)

The first order ordinary differential for the Hazard function of the truncated exponential distribution is given as;

$$h'(t) - h^{2}(t) + \lambda h(t) = 0$$
(128)

$$h(0) = \frac{1}{\lambda(1 - e^{-\frac{b}{\lambda}})}$$
(129)

To obtain the second order differential equation, differentiate equation (125) to obtain;

$$h''(t) = \frac{2}{\lambda^{3}} \left( e^{-\frac{t}{\lambda}} \right)^{3} \left( e^{-\frac{t}{\lambda}} - e^{-\frac{b}{\lambda}} \right)^{-3}$$
$$-\frac{2}{\lambda^{3}} \left( e^{-\frac{t}{\lambda}} \right)^{2} \left( e^{-\frac{t}{\lambda}} - e^{-\frac{b}{\lambda}} \right)^{-2} - \frac{1}{\lambda^{3}} \left( e^{-\frac{t}{\lambda}} \right)^{2} \left( e^{-\frac{t}{\lambda}} - e^{-\frac{b}{\lambda}} \right)^{-2}$$
$$+ \frac{1}{\lambda^{3}} \left( e^{-\frac{t}{\lambda}} \right) \left( e^{-\frac{t}{\lambda}} - e^{-\frac{b}{\lambda}} \right)^{-1}$$
(130)

The condition necessary for the existence of the equation is  $\lambda, b > 0, t \ge 0$ .

$$h''(t) = \frac{2}{\lambda^{3}} \left( e^{-\frac{t}{\lambda}} \right)^{3} \left( e^{-\frac{t}{\lambda}} - e^{-\frac{b}{\lambda}} \right)^{-3}$$

$$-\frac{3}{\lambda^{3}} \left( e^{-\frac{t}{\lambda}} \right)^{2} \left( e^{-\frac{t}{\lambda}} - e^{-\frac{b}{\lambda}} \right)^{-2} - \frac{1}{\lambda^{3}} \left( e^{-\frac{t}{\lambda}} \right) \left( e^{-\frac{t}{\lambda}} - e^{-\frac{b}{\lambda}} \right)^{-1}$$
(131)
$$h''(t) = \frac{\frac{2}{\lambda^{3}} \left( e^{-\frac{t}{\lambda}} \right)^{3}}{\left( e^{-\frac{t}{\lambda}} - e^{-\frac{b}{\lambda}} \right)^{3}} - \frac{\frac{3}{\lambda^{3}} \left( e^{-\frac{t}{\lambda}} \right)^{2}}{\left( e^{-\frac{t}{\lambda}} - e^{-\frac{b}{\lambda}} \right)^{2}} + \frac{\frac{1}{\lambda^{3}} \left( e^{-\frac{t}{\lambda}} \right)}{\left( e^{-\frac{t}{\lambda}} - e^{-\frac{b}{\lambda}} \right)^{3}}$$
(132)

$$h''(t) = 2h^{3}(t) - 3\lambda h^{2}(t) + \lambda^{2}h(t)$$
(133)

The second order ordinary differential for the Hazard function of the truncated exponential distribution is given as;

$$h''(t) - 2h^{3}(t) + 3\lambda h^{2}(t) - \lambda^{2}h(t) = 0$$
(134)

 $h'(0) = \frac{e^{-\frac{b}{\lambda}}}{\lambda^2 (1 - e^{-\frac{b}{\lambda}})^2}$ (135)

To obtain the third order differential equation, differentiate equation (131) to obtain;

$$h'''(t) = \frac{6}{\lambda^4} \left( e^{-\frac{t}{\lambda}} \right)^4 \left( e^{-\frac{t}{\lambda}} - e^{-\frac{b}{\lambda}} \right)^{-4} - \frac{6}{\lambda^4} \left( e^{-\frac{t}{\lambda}} \right)^3 \left( e^{-\frac{t}{\lambda}} - e^{-\frac{b}{\lambda}} \right)^{-3} - \frac{6}{\lambda^4} \left( e^{-\frac{t}{\lambda}} \right)^3 \left( e^{-\frac{t}{\lambda}} - e^{-\frac{b}{\lambda}} \right)^{-3} + \frac{6}{\lambda^4} \left( e^{-\frac{t}{\lambda}} \right)^2 \left( e^{-\frac{t}{\lambda}} - e^{-\frac{b}{\lambda}} \right)^{-2} + \frac{1}{\lambda^4} \left( e^{-\frac{t}{\lambda}} \right)^2 \left( e^{-\frac{t}{\lambda}} - e^{-\frac{b}{\lambda}} \right)^{-2} - \frac{1}{\lambda^4} \left( e^{-\frac{t}{\lambda}} \right) \left( e^{-\frac{t}{\lambda}} - e^{-\frac{b}{\lambda}} \right)^{-1}$$
(136)

The condition necessary for the existence of the equation is  $\lambda, b > 0, t \ge 0$ .

$$h'''(t) = \frac{6}{\lambda^4} \left( e^{-\frac{t}{\lambda}} \right)^4 \left( e^{-\frac{t}{\lambda}} - e^{-\frac{b}{\lambda}} \right)^{-4} - \frac{12}{\lambda^4} \left( e^{-\frac{t}{\lambda}} \right)^3 \left( e^{-\frac{t}{\lambda}} - e^{-\frac{b}{\lambda}} \right)^{-3} + \frac{7}{\lambda^4} \left( e^{-\frac{t}{\lambda}} \right)^2 \left( e^{-\frac{t}{\lambda}} - e^{-\frac{b}{\lambda}} \right)^{-2} - \frac{1}{\lambda^4} \left( e^{-\frac{t}{\lambda}} \right) \left( e^{-\frac{t}{\lambda}} - e^{-\frac{b}{\lambda}} \right)^{-1}$$
(137)

$$h'''(t) = 6h^4(t) - 12\lambda h^3(t) + 7\lambda^2 h^2(t) - \lambda^3 h(t)$$
(138)

The third order ordinary differential for the Hazard function of the truncated exponential distribution is given as;

$$h'''(t) - 6h^{4}(t) + 12\lambda h^{3}(t) - 7\lambda^{2}h^{2}(t) + \lambda^{3}h(t) = 0$$
(139)

$$h''(0) = \frac{e^{-\frac{b}{\lambda}}(1 - e^{-\frac{b}{\lambda}})}{\lambda^3 (1 - e^{-\frac{b}{\lambda}})^3}$$
(140)

## **REVERSED HAZARD FUNCTION**

The reversed hazard function (RSF) of the truncated exponential distribution is given by;

$$j(t) = \frac{e^{-\frac{t}{\lambda}}}{\lambda(1 - e^{-\frac{t}{\lambda}})}$$
(141)

The first order differential equation for the RHF can be obtained from the differentiation of equation (141);

$$j'(t) = \left\{ -\frac{\frac{1}{\lambda}e^{-\frac{t}{\lambda}}}{e^{-\frac{t}{\lambda}}} - \frac{\frac{1}{\lambda}e^{-\frac{t}{\lambda}}(1 - e^{-\frac{t}{\lambda}})^{-2}}{(1 - e^{-\frac{t}{\lambda}})^{-1}} \right\} j(t) \quad (142)$$

The condition necessary for the existence of the equation is  $\lambda, b > 0, t \ge 0.$ 

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$$j'(t) = -\left(\frac{1}{\lambda} + \frac{e^{-\frac{t}{\lambda}}}{\lambda(1 - e^{-\frac{t}{\lambda}})}\right)j(t)$$
(143)

$$j'(t) = -\left(\frac{1}{\lambda} + j(t)\right)j(t) \tag{144}$$

The first order ordinary differential for the reverse hazard function of the truncated exponential distribution is given  $2 \frac{1}{2} \frac{1}{2}$ 

as; 
$$\lambda j'(t) + \lambda j^{2}(t) + j(t) = 0$$
  
(145)  $j(1) = \frac{e^{-\frac{1}{\lambda}}}{\lambda(1 - e^{-\frac{1}{\lambda}})} = \frac{1}{\lambda(e^{\frac{1}{\lambda}} - 1)}$ 

(146) The ODEs of all the probability functions considered can be obtained for the particular values of the distribution. Several analytic, semi-analytic and numerical methods can be applied to obtain the solutions of the respective differential equations [29-32]. Also comparison with two or more solution methods is useful in understanding the link between ODEs and the probability distributions.

#### **III. CONCLUDING REMARKS**

In this work, differentiation was used to obtain some classes of ordinary differential equations for the probability density function (PDF), quantile function (QF), survival function (SF), inverse survival function (ISF) and reversed hazard function (RHF) of the exponential and truncated exponential distributions. The nature of exponential distribution prohibits the availability of the ODE of the hazard function. However, that was not the case with the truncated exponential distributions. In all, the parameters that define the respective distributions determine the nature of the respective ODEs and the range determines the existence of the ODEs.

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#### REFERENCES

- [1] A.W. Marshall and I. Olkin, "A multivariate exponential distribution", *J. Amer. Stat. Assoc.*, vol. 62, no. 317, pp. 30-44, 1967.
- [2] R.D. Gupta and D. Kundu, "Generalized exponential distribution: different method of estimations", J. Stat. Comput. Simul., vol. 69, no. 4, pp. 315-337, 2001.
- [3] S. Nadarajah and S. Kotz, "The beta exponential distribution", *Relia. Engine. Syst. Safety*, vol. 91, no. 6, pp. 689-697, 2006.
- [4] W. Barreto-Souza, A.H. Santos and G.M. Cordeiro, "The beta generalized exponential distribution", *J. Stat. Comput. Simul.*, vol. 80, no. 2, pp. 159-172, 2010.
- [5] D.R. Cox and N. Wermuth, "A note on the quadratic exponential binary distribution", *Biometrika*, vol. 81, no. 2, pp. 403-408, 1994.
- [6] W. Barreto-Souza and F. Cribari-Neto, "A generalization of the exponential-Poisson distribution", *Stat. Prob. Lett.*, vol. 79, no. 24, pp. 2493-2500, 2009.
- [7] A. Fatima and A. Roohi, "Extended Poisson exponential distribution", *Pak. J. Stat. Oper. Res.*, vol. 11, no. 3, pp. 361-375, 2015.

- [8] M.M. Ristić and N. Balakrishnan, "The gamma-exponentiated exponential distribution", J. Stat. Comput. Simul., vol. 82, no. 8, pp. 1191-1206, 2012.
- [9] O.I. Shittu, K.A. Adepoju and O.S. Yaya, "On the convoluted betaexponential distribution", J. Modern Math. Stat., vol. 6, no. 3-6, pp. 14-22, 2012.
- [10] R.D. Gupta and D. Kundu, "A new class of weighted exponential distributions", *Statistics*, vol. 43, no. 6, pp. 621-634, 2009.
- [11] T.A. Anake, P.E. Oguntunde and O.A. Odetunmibi, "On a fractional beta-exponential distribution", *Int. J. Math. Comp.*, vol. 26, no. 1, pp. 26-34, 2015.
- [12] J.A. Fioruci, B. Yiqi, F. Louzada and V.G. Cancho, "The exponential Poisson logarithmic distribution", *Comm. Stat. Theo. Meth.*, vol. 45, no. 9, pp. 2556-2575, 2016.
- [13] J.A. Rodrigues and A.P. Silva, "The exponentiated Kumaraswamyexponential distribution", *British J. Appl. Sci. Tech.*, vol. 10, no. 5, pp. 1-12, 2015.
- [14] P.E. Oguntunde, O.S. Balogun, H.I. Okagbue and S.A. Bishop, "The Weibull-exponential distribution: Its properties and applications", J. Appl. Sciences, vol. 15, no. 11, pp. 1305-1311, 2015.
- [15] P.E. Oguntunde, A.O. Adejumo, H.I. Okagbue and M.K. Rastogi, "Statistical properties and applications of a new Lindley exponential distribution", *Gazi Uni. J. Science*, vol. 29, no. 4, pp. 831-838, 2016.
- [16] P.E. Oguntunde, E.A. Owoloko and O.S. Balogun, "On a new weighted exponential distribution: theory and application", *Asian J. Appl. Sci.*, vol. 9, no. 1, pp. 1-12, 2016.
- [17] L.J. Bain and D.L. Weeks, "A note on the truncated exponential distribution", Ann. Math. Stat., vol. 35, no. 3, pp. 1366-1367, 1964.
- [18] G. Lominashvili and M. Patsatsia, "On the estimation of a maximum likelihood of truncated exponential distributions", *Bul. Georgian Nat. Acad. Sci.*, vol. 7, no. 1, pp. 21-24, 2013.
- [19] M.C. Korkmaz and A.I. Genç, "Two-Sided Generalized Exponential Distribution", *Comm. Stat. Theo. Meth.*, vol. 44, no. 23, pp. 5049-5070, 2015.
- [20] R.D. Gupta and D. Kundu, "Generalized exponential distributions", *Aust. New Zeal. J. Stat.*, vol. 41, no. 2, pp. 173-188, 1999.
- [21] A.B. Rad, "Thermo-Elastic Analysis of Non-Uniform Functionally Graded Circular Plate Resting on a Gradient Elastic Foundation", J. Solid Mech., vol. 9, no. 1, pp. 63-85, 2017.
- [22] Y. Wang, K. Liu and Y. Lou, "An age-structured within-host HIV model with T-cell competition", *Nonl. Analy. Real World Appl.*, 38, 1-20, 2017.
- [23] I. Popa, C. Nechita and A. Hofgaard, "Stand structure, recruitment and growth dynamics in mixed subalpine spruce and Swiss stone pine forests in the Eastern Carpathians", *Science Total Environ.*, vol. 598, pp. 1050-1057, 2017.
- [24] N.L. Johnson, S. Kotz and N. Balakrishnan, Continuous univariate distributions, Wiley New York. ISBN: 0-471-58495-9, 1994.
- [25] W.P. Elderton, Frequency curves and correlation, Charles and Edwin Layton. London, 1906.
- [26] H. Rinne, Location scale distributions, linear estimation and probability plotting using MATLAB, 2010.
- [27] N. Balakrishnan and C.D. Lai, Continuous bivariate distributions, 2nd edition, Springer New York, London, 2009.
- [28] N.L. Johnson, S. Kotz and N. Balakrishnan, Continuous Univariate Distributions, Volume 2. 2nd edition, Wiley, 1995.
- [29] A.A. Opanuga, S.O. Edeki, H.I. Okagbue and G.O. Akinlabi, "Numerical solution of two-point boundary value problems via differential transform method", *Global J. Pure Appl. Math.*, vol. 11, no. 2, pp. 801-806, 2015.
- [30] A.A Opanuga, O.O. Agboola and H.I. Okagbue, "Approximate solution of multipoint boundary value problems", J. Engine. Appl. Sci., vol. 10, no. 4, pp. 85-89, 2015.
- [31] A.A. Opanuga, O.O. Agboola, H.I. Okagbue and J.G. Oghonyon, "Solution of differential equations by three semi-analytical techniques", *Int. J. Appl. Engine. Res.*, vol. 10, no. 18, pp. 39168-39174, 2015.
- [32] A.A. Opanuga, S.O. Edeki, H.I. Okagbue, S.A. Adeosun and M.E. Adeosun, "Some Methods of Numerical Solutions of Singular System of Transistor Circuits", J. *Comp. Theo. Nanosci.*, vol. 12, no. 10, pp. 3285-3289, 2015.